### Uncertainty Reduction in Atmospheric Composition Models by Chemical Data Assimilation

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Aug. 4, 2011. IFIP UQ Workshop, Boulder, CO



## Information feedback loops between CTMs and observations: data assimilation and targeted meas.



### What is data assimilation?

- The fusion of information from:
- 1. prior knowledge,
- 2. imperfect model predictions, and
- 3. sparse and noisy data,

to obtain a consistent description of the state of a physical system, such as the atmosphere.



Lars Isaksen (http://www.ecmwf.int)



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### Source of information #1: The prior encapsulates our *current knowledge* of the state

- The background (prior) probability density:  $\mathcal{P}^{b}(\boldsymbol{x})$
- ► The current best estimate: apriori (background) state **x**<sup>b</sup>.
- Typical assumption on random background errors

$$arepsilon^{\mathrm{b}} = \mathbf{X}^{\mathrm{b}} - \mathcal{S}(\mathbf{X}^{\mathrm{true}}) \in \mathcal{N}\left(\mathbf{0}, \mathbf{B}
ight)$$
 .

With many nonlinear models the normality assumption is difficult to justify, but is nevertheless widely used because of its convenience.





Source of information #2: The model encapsulates our knowledge about physical and chemical laws that govern the evolution of the system

▶ The model evolves an initial state  $\mathbf{x}_0 \in \mathbb{R}^n$  to future times

$$\mathbf{x}_{i} = \mathcal{M}_{t_{0} \rightarrow t_{i}} (\mathbf{x}_{0})$$
.

- ▶ Typical size of chemical transport models:  $n \in O(10^7)$  variables.
- The model is imperfect

$$\mathcal{S}\left(\mathbf{x}_{i}^{\text{true}}\right) = \mathcal{M}_{t_{i-1} \to t_{i}} \cdot \mathcal{S}\left(\mathbf{x}_{i-1}^{\text{true}}\right) - \eta_{i},$$

where  $\eta_i$  is the model error in step *i*.





### Source of information #3: The observations are sparse and noisy snapshots of reality

• Measurements  $\mathbf{y}_i \in \mathbb{R}^m$  ( $m \ll n$ ) taken at times  $t_1, \ldots, t_N$ 

 $\mathbf{y}_{i} = \mathcal{H}^{t}\left(\mathbf{x}_{i}^{\text{true}}\right) - \varepsilon_{i}^{\text{instrument}} = \mathcal{H}\left(\mathcal{S}(\mathbf{x}_{i}^{\text{true}})\right) - \varepsilon_{i}^{\text{obs}}, \quad i = 1, \cdots, N.$ 

- Observation operators
  - $\mathcal{H}^t$  : physical space  $\rightarrow$  observation space, while
  - $\mathcal{H}$  : the model space  $\rightarrow$  observation space.
- The observation error

$$\varepsilon_{i}^{\text{obs}} = \underbrace{\varepsilon_{i}^{\text{instrument}}}_{\text{instrument error}} + \underbrace{\mathcal{H}\left(\mathcal{S}(\mathbf{x}_{i}^{\text{true}})\right) - \mathcal{H}^{\text{t}}\left(\mathbf{x}_{i}^{\text{true}}\right)}_{\text{representativeness error}}$$

Typical assumptions:

$$\varepsilon_{i}^{\text{obs}} \in \mathcal{N}(\mathbf{0}, \mathbf{R}_{i})$$
;  $\varepsilon_{i}^{\text{obs}}$ ,  $\varepsilon_{j}^{\text{obs}}$  independent for  $t_{i} \neq t_{j}$ .



Data assimilation. General view of data assimilation. Sources of information [9/21] Aug. 4, 2011. IFIP UQ Workshop, Boulder, CO. (http://csl.cs.vt.edu)



### Result of data assimilation: The analysis encapsulates our *enhanced knowledge* of the state

► The analysis (posterior) probability density  $\mathcal{P}^{a}(\mathbf{x})$ :

Bayes: 
$$\mathcal{P}^{a}(\mathbf{x}) = \mathcal{P}(\mathbf{x}|\mathbf{y}) = \frac{\mathcal{P}(\mathbf{y}|\mathbf{x}) \cdot \mathcal{P}^{b}(\mathbf{x})}{\mathcal{P}(\mathbf{y})}$$

- ► The best state estimate **x**<sup>a</sup> is called the aposteriori, or the *analysis*.
- Analysis estimation errors  $\varepsilon^{a} = \mathbf{x}^{a} \mathcal{S}(\mathbf{x}^{true})$  characterized by
  - analysis mean error (bias)  $\beta^{a} = \mathbb{E}^{a} [\varepsilon^{a}]$
  - analysis error covariance matrix

$$\mathbf{A} = \mathbb{E}^{\mathrm{a}}\left[ \left( \varepsilon^{\mathrm{a}} - \beta^{\mathrm{a}} \right) \left( \varepsilon^{\mathrm{a}} - \beta^{\mathrm{a}} \right)^{\mathsf{T}} \right] \in \mathbb{R}^{n \times n}$$

 In the Gaussian, linear case, Bayes posterior admits an analytical solution by Kalman filter formulas





#### Extended Kalman filter

- The observations are considered successively at times  $t_1, \dots, t_N$ .
- ► The background state at *t<sub>i</sub>* given by the model forecast:

$$\mathbf{x}_{i}^{\mathrm{b}} \equiv \mathbf{x}_{i}^{\mathrm{f}} = \mathcal{M}_{t_{i-1} \to t_{i}} \cdot \mathbf{x}_{i-1}^{\mathrm{a}}$$

Model is imperfect, but is assumed unbiased

$$\eta_i \in \mathcal{N}(\mathbf{0}, \mathbf{Q}_i)$$

 Model error η<sub>i</sub> and solution error ε<sup>a</sup><sub>i-1</sub> are assumed independent; solution error small, propagated by linearized model M = M'(x)

$$\mathcal{O}(n^3): \quad \mathbf{B}_i \equiv \mathbf{P}_i^{\mathrm{f}} = \mathbf{M}_{t_{i-1} \rightarrow t_i} \, \mathbf{P}_{i-1}^{\mathrm{a}} \, \mathbf{M}_{t_i \rightarrow t_{i-1}}^T + \mathbf{Q}_i \, .$$

• EKF analysis uses  $\mathbf{H}_i = \mathcal{H}'(\mathbf{x}_i^{\mathrm{f}})$ :

$$\begin{aligned} \mathcal{O}(nm) : & \mathbf{x}_i^{\mathrm{a}} = \mathbf{x}_i^{\mathrm{f}} + \mathbf{K}_i \, \left( \mathbf{y}_i - \mathcal{H}(\mathbf{x}_i^{\mathrm{f}}) \right) \\ \mathcal{O}(nm^2 + n^2m + m^3) : & \mathbf{K}_i = \mathbf{P}_i^{\mathrm{f}} \mathbf{H}_i^{\mathrm{T}} \left( \mathbf{H}_i \, \mathbf{P}_i^{\mathrm{f}} \, \mathbf{H}_i^{\mathrm{T}} + \mathbf{R}_i \right)^{-1} \\ \mathcal{O}(n^2m + n^3) : & \mathbf{A}_i \equiv \mathbf{P}_i^{\mathrm{a}} = \left( \mathbf{I} - \mathbf{K}_i \, \mathbf{H}_i \right) \, \mathbf{P}_i^{\mathrm{f}}. \end{aligned}$$



Data assimilation. The Bayesian framework. [11/21] Aug. 4, 2011. IFIP UQ Workshop, Boulder, CO. (http://csl.cs.vt.edu)



#### Practical Kalman filter methods

- EKF is not practical for very large systems
- Suboptimal KF approximate the covariance matrices e.g.,

$$\mathbf{B}_{(\ell),(k)} = \sigma_{(\ell)} \, \sigma_{(k)} \, \exp\left(\operatorname{distance}\{\operatorname{gridpoint}(\ell), \operatorname{gridpoint}(k)\}^2 / L^2\right)$$

Ensemble Kalman filters (EnKF) use a Monte-Carlo approach

$$\mathbf{x}_{i}^{\mathrm{f}}[\boldsymbol{e}] = \mathcal{M}_{t_{i-1} \to t_{i}} \left( \mathbf{x}_{i-1}^{\mathrm{a}}[\boldsymbol{e}] \right) + \underbrace{\eta_{i}[\boldsymbol{e}]}_{\mathrm{model \, error}}, \quad \boldsymbol{e} = 1, \dots, \boldsymbol{E}$$

$$\mathbf{x}_{i}^{a}[\boldsymbol{e}] = \mathbf{x}_{i}^{f}[\boldsymbol{e}] + \mathbf{K}_{i}\left(\mathbf{y}_{i} + \varepsilon_{i}^{\text{obs}}[\boldsymbol{e}] - \mathcal{H}_{i}(\mathbf{x}_{i}^{f}[\boldsymbol{e}])\right), \quad \boldsymbol{e} = 1, \dots, \boldsymbol{E}.$$

- Error covariances  $\mathbf{P}_{i}^{f}$ ,  $\mathbf{P}_{i}^{a}$  estimated from statistical samples
- EnKF issues: rank-deficiency of the estimated P<sup>f</sup><sub>i</sub>
- EnKF strengths: capture non-linear dynamics, doesn't need TLM, ADJ, accounts for model errors, almost ideally parallelizable





#### Maximum aposteriori estimator

Maximum aposteriori estimator (MAP) defined by

$$\mathbf{x}^{a} = \arg \max_{\mathbf{x}} \, \mathcal{P}^{a}(\mathbf{x}) = \arg \min_{\mathbf{x}} \, \mathcal{J}(\mathbf{x}) \,, \quad \mathcal{J}(\mathbf{x}) = - \ln \, \mathcal{P}^{a}(\mathbf{x}) \,.$$

Using Bayes and assumptions for background, observation errors:

$$\begin{aligned} \mathcal{J}(\mathbf{x}) &= -\ln \mathcal{P}^{a}(\mathbf{x}) = -\ln \mathcal{P}^{b}\left(\mathbf{x}\right) - \ln \mathcal{P}\left(\mathbf{y}|\mathbf{x}\right) + \text{const} \\ &\doteq \frac{1}{2} \left(\mathbf{x} - \mathbf{x}^{b}\right)^{T} \mathbf{B}^{-1} \left(\mathbf{x} - \mathbf{x}^{b}\right) + \frac{1}{2} \left(\mathcal{H}\left(\mathbf{x}\right) - \mathbf{y}\right)^{T} \mathbf{R}^{-1} \left(\mathcal{H}\left(\mathbf{x}\right) - \mathbf{y}\right) \end{aligned}$$

Optimization by gradient-based numerical procedure

$$\nabla_{\boldsymbol{x}} \mathcal{J} \left( \boldsymbol{x}^a \right) = \boldsymbol{\mathsf{B}}^{-1} \ \left( \boldsymbol{x}^a - \boldsymbol{x}^b \right) + \boldsymbol{\mathsf{H}}^{\mathsf{T}} \, \boldsymbol{\mathsf{R}} \ \left( \mathcal{H} (\boldsymbol{x}^a) - \boldsymbol{y} \right) \, ; \quad \boldsymbol{\mathsf{H}} = \mathcal{H} (\boldsymbol{x}^b) \, .$$

Hessian of cost function approximates inverse analysis covariance

$$abla^2_{\mathbf{x},\mathbf{x}}\mathcal{J} = \mathbf{B}^{-1} + \mathbf{H}^T \, \mathbf{R}^{-1} \, \mathbf{H} \approx \mathbf{A}^{-1}$$





### Four dimensional variational data assimilation (4D-Var) I

- ► All observations at all times t<sub>1</sub>, · · · , t<sub>N</sub> are considered simultaneously
- The control variables (parameters p, initial conditions x<sub>0</sub>, boundary conditions, etc) uniquely determine the state of the system at all future times
- 4D-Var MAP estimate via model-constrained optimization problem

$$\begin{aligned} \mathcal{J}\left(\mathbf{x}_{0}\right) &= \frac{1}{2} \left\|\mathbf{x}_{0} - \mathbf{x}_{0}^{b}\right\|_{\mathbf{B}_{0}^{-1}}^{2} + \frac{1}{2} \sum_{i=1}^{N} \left\|\mathcal{H}(\mathbf{x}_{i}) - \mathbf{y}_{i}\right\|_{\mathbf{R}_{i}^{-1}}^{2} \\ \mathbf{x}_{0}^{a} &= \arg\min \mathcal{J}\left(\mathbf{x}_{0}\right) \\ &\text{subject to: } \mathbf{x}_{i} = \mathcal{M}_{t_{0} \to t_{i}}\left(\mathbf{x}_{0}\right), \quad i = 1, \cdots, N \end{aligned}$$

Formulation can be easily extended to other model parameters





### Four dimensional variational data assimilation (4D-Var) II

- The large scale optimization problem is solved in a reduced space using a gradient-based technique.
- The 4D-Var gradient reads

$$\nabla \mathcal{J}_{\mathbf{x}_0}\left(\mathbf{x}_0\right) = \mathbf{B}_0^{-1} \left(\mathbf{x}_0 - \mathbf{x}_0^{\mathrm{b}}\right) + \sum_{i=1}^{N} \left(\frac{\partial \mathbf{x}_i}{\partial \mathbf{x}_0}\right)^T \mathbf{H}_i^T \mathbf{R}_i^{-1} \left(\mathcal{H}(\mathbf{x}_i) - \mathbf{y}_i\right)$$

- ► Needs linearized observation operators  $\mathbf{H}_i = \mathcal{H}'(\mathbf{x}_i)$
- ► Needs the transposed sensitivity matrix  $(\partial \mathbf{x}_i / \partial \mathbf{x}_0)^T \in \mathbb{R}^{n \times n}$
- Adjoint models efficiently compute the transposed sensitivity matrix times vector products
- The construction of an adjoint model is a nontrivial task.





## Correct models of background errors are of great importance for data assimilation

- Background error representation determines the spread of information, and impacts the assimilation results
- Needs: high rank, capture dynamic dependencies, efficient computations
- Traditionally estimated empirically (NMC, Hollingsworth-Lonnberg)
- 1. Tensor products of 1d correlations, decreasing with distance (Singh et al, 2010)
- Multilateral AR model of background errors based on "monotonic TLM discretizations" (Constantinescu et al 2007)
- 3. Hybrid methods in the context of 4D-Var (Cheng et al, 2007)



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#### What is the effect of mis-specification of inputs?

(Daescu, 2008) Consider a *verification functional*  $\Psi(\mathbf{x}_{v}^{a})$  defined on the optimal solution at a future time  $t_{v}$ .  $\Psi$  is a measure of the forecast error. What is the impact of small errors in the specification of covariances, background, and observation data?

$$\begin{aligned} \nabla_{\mathbf{y}_{i}} \Psi &= \mathbf{R}_{i}^{-1} \mathbf{H}_{i} \mathbf{M}_{t_{0} \to t_{i}} \left( \left( \nabla_{\mathbf{x}_{0}, \mathbf{x}_{0}}^{2} \mathcal{J} \right)^{-1} \nabla_{\mathbf{x}_{0}} \Psi \right) \\ \nabla_{\mathbf{R}_{i}(:)} \Psi &= \left( \mathbf{R}_{i}^{-1} (\mathcal{H}(\mathbf{x}_{i}^{a}) - \mathbf{y}_{i}) \right) \otimes \nabla_{\mathbf{y}_{i}} \Psi \\ \nabla_{\mathbf{x}^{b}} \Psi &= \mathbf{B}_{0}^{-1} \left( \left( \nabla_{\mathbf{x}_{0}, \mathbf{x}_{0}}^{2} \mathcal{J} \right)^{-1} \nabla_{\mathbf{x}_{0}} \Psi \right) \\ \nabla_{\mathbf{B}_{0}(:)} \Psi &= \left( \mathbf{B}_{0}^{-1} (\mathbf{x}_{0}^{a} - \mathbf{x}_{0}^{b}) \right) \otimes \nabla_{\mathbf{x}^{b}} \Psi \end{aligned}$$





#### General framework for sensitivity analysis Forward model equations link parameters and solutions:

$$\begin{split} \mathcal{F}(\mathbf{x},\theta) &= \mathbf{0} \in \mathcal{H}_F \ . \quad \mathcal{H}_F \text{= model constraint space, Hilbert: } \langle \cdot, \cdot \rangle_{\mathcal{H}_F} \\ \mathbf{x} \in \mathcal{H}_{\mathbf{x}} \ . \quad \mathcal{H}_{\mathbf{x}} \text{= model state space, Hilbert: } \langle \cdot, \cdot \rangle_{\mathcal{H}_{\mathbf{x}}} \ , \\ \theta \in \mathcal{H}_{\theta} \ . \quad \mathcal{H}_{\theta} \text{= parameter space, Hilbert: } \langle \cdot, \cdot \rangle_{\mathcal{H}_{\theta}} \ . \end{split}$$

The response functional (QoI) associates a real value to each state

$$\mathcal{J}(\boldsymbol{x}): \ \mathcal{H}_{\boldsymbol{x}} \ \longrightarrow \ \mathbb{R} \quad \left( e.g., \ \mathcal{J}(\boldsymbol{x}) = \frac{1}{2} \ \|\mathcal{H}(\boldsymbol{x}) - \boldsymbol{y}\|_{\boldsymbol{\mathsf{R}}^{-1}}^2 \right)$$

Assumptions:

- 1.  $\mathcal{F}, \mathcal{J}$  are continuously Frèchet differentiable.
- 2.  $\mathcal{F}_{\mathbf{x}}$  has a continuous linear inverse mapping. By IFT a Frèchet differentiable model solution operator  $\mathbf{x} = \mathcal{M}(\theta)$  exists locally

$$\mathcal{M}: \mathcal{H}_{\theta} \to \mathcal{H}_{\mathbf{x}}; \quad \mathbf{x} = \mathcal{M}(\theta); \quad \mathcal{M}'(\theta) = -\mathcal{F}_{\mathbf{x}}^{-1}(\mathbf{x}, \theta) \cdot \mathcal{F}_{\theta}(\mathbf{x}, \theta).$$





Formulation of the inverse problem as a model-constrained optimization problem Find the optimal vector of parameters  $\theta_{opt}$  such that:

Comments.

1. The cost function depends implicitly on the parameters:

$$\mathcal{J}(\mathbf{X}) = \mathcal{J}\left(\mathcal{M}(\theta)\right) = \left(\mathcal{J} \circ \mathcal{M}\right) \ (\theta) \ .$$

2. Gradient-based optimization techniques require

$$abla_{ heta}\mathcal{J}=\mathcal{M}^{\prime*}( heta)\cdot
abla_{\mathbf{X}}\mathcal{J}$$

3. Difficulty: model solution operator is only defined implicitly.





#### Direct (forward) vs. adjoint sensitivity analysis

1. Tangent linear model is obtained by Frèchet differentiation

$$(\textit{TLM}): \quad \mathcal{F}_{\theta}(\theta, \mathbf{X}) \cdot \delta\theta + \mathcal{F}_{\mathbf{X}}(\theta, \mathbf{X}) \cdot \delta\mathbf{X} = \mathbf{0} \in \mathcal{H}_{\textit{F}} \; .$$

$$\delta \mathcal{J} = \langle \nabla_{\mathbf{x}} \mathcal{J}, \delta \mathbf{x} \rangle_{\mathcal{H}_{\mathbf{x}}} = \langle \nabla_{\theta} \mathcal{J}, \delta \theta \rangle_{\mathcal{H}_{\theta}} .$$

Comment.  $\nabla_{\mathbf{x}} \mathcal{J}$  by direct differentiation. One TLM solution provides one inner product . To find the entire gradient  $\nabla_{\theta} \mathcal{J}$  ...





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Comment.  $\nabla_{\mathbf{x}} \mathcal{J}$  by direct differentiation. One TLM solution provides one inner product . To find the entire gradient  $\nabla_{\theta} \mathcal{J}$  ... 2. *Adjoint model* obtained using duality:

$$\begin{split} &(\lambda \in \mathcal{H}_{F}^{*} \equiv \mathcal{H}_{F}) \iff \langle \lambda, \mathcal{F}_{\mathbf{x}} \cdot \delta \mathbf{x} \rangle_{\mathcal{H}_{F}} + \langle \lambda, \mathcal{F}_{\theta} \cdot \delta \theta \rangle_{\mathcal{H}_{F}} = \mathbf{0} \in \mathbb{R} \\ &(\textit{by adjoint}) \iff \langle \mathcal{F}_{\mathbf{x}}^{*} \cdot \lambda, \delta \mathbf{x} \rangle_{\mathcal{H}_{\mathbf{x}}} + \langle \mathcal{F}_{\theta}^{*} \cdot \lambda, \delta \theta \rangle_{\mathcal{H}_{\theta}} = \mathbf{0} \in \mathbb{R} \,. \\ &ADJ : \quad \mathcal{F}_{\mathbf{x}}^{*} \cdot \lambda = -\nabla_{\mathbf{x}} \mathcal{J}(\mathbf{x}) \,. \\ &\langle \nabla_{\mathbf{x}} \mathcal{J}(\mathbf{x}), \delta \mathbf{x} \rangle_{\mathcal{H}_{\mathbf{x}}} = \langle (\mathcal{F}_{\theta})^{*} \cdot \lambda, \delta \theta \rangle_{\mathcal{H}_{\theta}} = \langle \nabla_{\theta} \mathcal{J}, \delta \theta \rangle_{\mathcal{H}_{\theta}} = \delta \mathcal{J} \,. \end{split}$$

Comment. Adjoint model does not depend on the particular perturbations  $\delta\theta$ ,  $\delta \mathbf{x}$ , and needs to be solved *only once*.





Continuous and discrete adjoints of mass balance equations lead to different computational models

$$\nabla_{\mathbf{y}^{0}} \psi = \dots + \sum_{k=1}^{N} \left( \partial \mathbf{y}^{k} / \partial \mathbf{y}^{0} \right)^{\mathbf{T}} \left( \mathbf{H}^{k} \right)^{\mathbf{T}} \mathbf{R}_{k}^{-1} \left( \mathbf{H}^{k} \mathbf{y}^{k} - \mathbf{z}_{obs}^{k} \right)$$





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# Discrete adjoints of advection numerical schemes can become inconsistent with the adjoint PDE



Change of forward scheme pattern:

- Change of upwinding
- Sources/sinks
- Inflow boundaries scheme Example: 3<sup>rd</sup> order upwind FD



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Discrete Runge-Kutta adjoints can be regarded as "numerical methods" applied to the adjoint ODE

RK Method  

$$\mathbf{y}^{\mathbf{n}+1} = \mathbf{y}^{\mathbf{n}} + h \sum_{i=1}^{s} b_i \mathbf{f}(\mathbf{Y}^i),$$
  
 $\mathbf{Y}^i = \mathbf{y}^{\mathbf{n}} + h \sum_{i=1}^{s} a_{i,j} \mathbf{f}(\mathbf{Y}^j)$  Discrete RK Adjoint  
*[Hager, 2000]*  
 $\lambda^{\mathbf{n}} = \lambda^{\mathbf{n}+1} + \sum_{i=1}^{s} \theta^i$   
 $\theta^i = h \mathbf{J}^{\mathbf{T}}(\mathbf{Y}^i) \cdot \left[ b_i \lambda^{\mathbf{n}+1} + \sum_{j=1}^{s} a_{j,i} \theta^j \right]$ 



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#### Discrete Runge-Kutta adjoints: error analysis

*Local error analysis:* The discrete adjoint of RK method of order p **is an order p** discretization of the adjoint equation. [Sandu, 2005]. This:

- works for both explicit and implicit methods
- true for arbitrary orders p

*Global error analysis:* The discrete adjoint (of a RK method convergent with order p) **converges with order p** to the solution of the adjoint ODE. [Sandu, 2005] The analysis accounts for:

- 1. the truncation error at each step, and
- *2. the different trajectories about which the continuous and the discrete adjoints are defined*

Stiff case: Consider a stiffly accurate Runge Kutta method **of order p** with invertible coefficient matrix A. The discrete adjoint provides:

- an **order p** discretization of the adjoint of **nonstiff variable**
- an order min(p,q+1,r+1) of the adjoint of stiff variable [Sandu, 2005]



Properties of discrete adjoint LMM

- 1. For fixed step sizes
  - the discrete adjoint starting and ending steps, in general, are not consistent approximations of the adjoint ODE
  - the adjoint LMM is (at least) first order consistent with the adjoint ODE
- 2. For **variable step sizes** the adjoint LMM is not a consistent discretization of the adjoint ODE
- 3. The discrete **adjoint variable at the initial time** is an order p approximation of the continuous adjoint, where p is the order of the (forward) LMM method. *[Sandu, 2007]*



### Uncertainty quantification using polynomial chaos and the STEM model







Data assimilation. The Bayesian framework. UQ/UA for STEM [23/25] Aug. 4, 2011. IFIP UQ Workshop, Boulder, CO. (http://csl.cs.vt.edu)



#### Uncertainty apportionment with the STEM model



Figure: Top: New York. Bottom: Boston. 48 hrs ozone mean, standard deviation, and uncertainty (variance) apportionment.





### Quantification of the probability of non-compliance with the NAAQS ozone maximum admissible levels



**Figure:** Boston 8hrs average ozone PDF shows a 68% probability of exceeding the maximum admissible level of 75 ppbv.





## Ensemble-based chemical data assimilation is an alternative to variational techniques



## The Ensemble Kalman Filter (EnKF) popular in NWP but not extensively used before with CTMs



[Constantinescu et al., 2007]

Ozonesonde S2 (18 EDT, July 20, 2004)

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## Ground level ozone at 2pm EDT, July 20, 2004, better matches observations after LEnKF data assimilation

Observations: circles, color coded by O<sub>3</sub> mixing ratio

Forecast (R<sup>2</sup>=0.24/0.28)

Analysis (R<sup>2</sup>=0.88/0.32)

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## The use of adjoints in large scale simulations: atmospheric chemical transport models



## Adjoint sensitivity analysis of non-attainment metrics can help guide policy decisions



## STEM: Assimilation adjusts O<sub>3</sub> predictions considerably at 4pm EDT on July 20, 2004

Observations: circles, color coded by O<sub>3</sub> mixing ratio



### Assimilation of elevated observations for July 20, 2004

NOAA P3 flight observations

Ozonesonde observations (Rhode Island)





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## The inversion procedure can be extended to emissions, boundary conditions, etc.

Texas: 4am CST July 16 to 8pm CST on July 17, 2004.



## Smallest Hessian eigenvalues (vectors) approximate the principal aposteriori error components

$$\left(\nabla^{2}_{y^{0},y^{0}}\Psi\right)^{-1} \approx \operatorname{cov}(y^{\operatorname{opt}})$$



(b) East view



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#### Assimilation of TES ozone column observations, August 2006. Lobatto-IIIC integrates stiff chemistry.









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## Quality of TES ozone column data assimilation results for several methods (August 1-15, 2006)



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Dynamic integration of chemical data and atmospheric models is an important, growing field

- the tools needed for 4D-Var chemical data assimilation are in place:
  - (adjoints for stiff systems, aerosols, transport; singular vectors, parallelization and multi-level checkpointing schemes, models of background errors)
- all algorithms are on a solid theoretical basis
- the ensemble filter methods show promise
- STEM, CMAQ, GEOS-CHEM have been endowed with data assimilation capabilities
- the tool strengths have been demonstrated using real (field campaign) data; ambitious science projects are ongoing



